ES1004 Econometrics by Example

Lecture 13: Economic Forecasting

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Gujarati textbook, second edition [chapter 16]

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Time Series Econometrics

- stationary and nonstationary time series
- cointegration and error correction models
- asset price volatility: the ARCH and GARCH models
- economic forecasting
 - previous course on time series econometrics

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Economic Forecasts

- Economics [GDP, unemployment, consumption, investment, interest rates]
- financial asset management [asset returns, exchange rates and commodity prices]
- financial risk management [asset return volatility]
- marketing [response of sales to different marketing schemes]
- business and government [revenue forecasts]
- crisis management [probabilities of default, currency devaluations]



Econometric Models

- based on past and current information, the objective of forecasting is to provide quantitative estimate(s) of the likelihood of the future course of the object of interest
 - e.g. personal consumption expenditure
- we develop econometric models and use one or more methods of forecasting its future course





Methods of Forecasting

- there are several methods of forecasting
- we will consider three prominent methods of forecasting in this chapter
 - regression models,
 - autoregressive integrated moving average (ARIMA) models [Box–Jenkins (BJ) methodology]
 - vector autoregression (VAR) models (Sims)





Example: Consumption Function

- forecasting is probably the most important purpose of estimating regression models
- consider the following bivariate regression

$$PCE_t = \beta_1 + \beta_2 PDI_t + u_t$$

where

PCE = per capita personal consumption expenditure,

PDI = per capita personal disposable





Example: Consumption Function

$$PCE_t = \beta_1 + \beta_2 PDI_t + u_t$$

- the slope coefficient represents the MPC
- to estimate this regression, we obtained aggregate data on these variables for the US for 1960–2008
- see Table16.1 on Piazza





Holdover Sample

- to estimate the consumption function, initially we use the observations from 1960–2004
- save the last four observations
 - called the holdover sample
 - to evaluate the performance of the estimated model



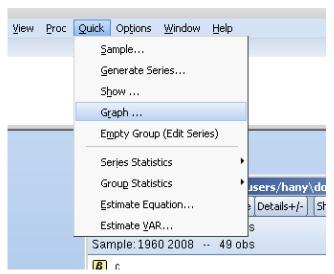


Plot the Data

We first plot the data to get some idea of the nature of the relationship between the two variables (Figure 16.1).

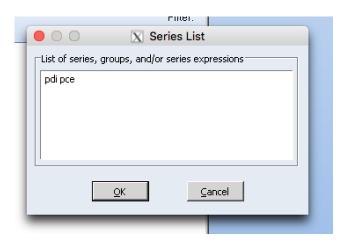






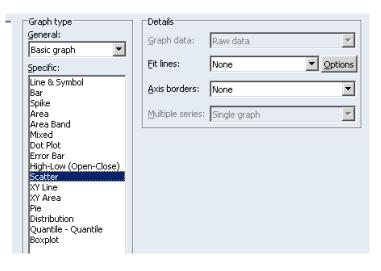




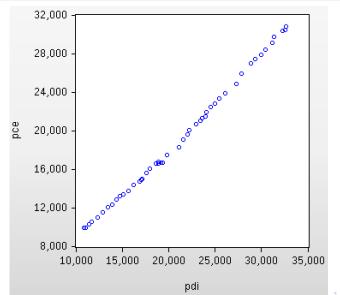
















Example: Estimation

- we plot the data to get some idea of the nature of the relationship between the two variables
- the figure shows almost a linear relationship between PCE and PDI
- we fitting a linear regression model to the data





Estimates of the consumption function, 1960–2004

	X Equation Estimation			
Specifica	tion Options .			
Equa	Equation specification Dependent variable followed by list of regressors including ARMA			
	and PDL terms, OR an explicit equation like Y=c(1)+c(2)*X.			
pce	c pdi			
L				
Estim	ation settings			
<u>M</u> eth	od: LS - Least Squares (NLS and ARMA)			
Samp	le: 1960 2004			





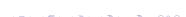
Estimates of the consumption function, 1960–2004

Dependent Variable: PCE Method: Least Squares Date: 02/12/17 Time: 23:48 Sample: 1960 2004

Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C PDI	-1083.978 0.953768	193.9579 0.009233	-5.588729 103.2981	0.0000 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.995986 0.995893 353.4907 5373095. -326.8829 10670.51 0.000000	Mean depend S.D. depende Akaike info cr Schwarz crite Hannan-Quin Durbin-Watsd	ent var iterion rion in criter.	18197.91 5515.914 14.61702 14.69731 14.64695 0.299775





Estimated Consumption Function

$$P\hat{C}E_t = -1083.978 + 0.9537PDI_t$$

- we can use this regression model to forecast the future value(s) PCE
- suppose we want to find out $E(PCE_{2005}|PDI_{2005})$ [$PDI_{2005} = \$31, 318$]





Point & Interval Forecasts

- in point forecasts we provide a single value for each forecast period
- in interval forecasts we obtain a range, or an interval, that will include the realized value with some probability
- the interval forecast provides a margin of uncertainty about the point forecast





ex post & ex ante

- estimation period: we have data on all the variables in the model
- ex post forecast period: we also know the values of the regressand and regressors
 - the holdover period used to get some idea about the performance of the fitted model
- ex ante forecast we estimate the values of the depend variable beyond the estimation period but we may not know the values of the regressors with certainty





ex post & ex ante

Estimation period	Ex-post forecast period	Ex-ante forecast period
1960-2004	2005-2008	2009 forward



Conditional & Unconditional Forecasts

- conditional forecasts: we forecast the variable of interest conditional on the assumed values of the regressors
- recall that all along we have conducted our regression analysis, conditional on the given values of the regressors.
- this type of conditional forecasting is also known as scenario analysis or contingency analysis





Conditional & Unconditional Forecasts

- in unconditional forecasts, we know the values of the regressors with certainty instead of picking some arbitrary values of them, as in conditional forecasting
- of course, that is a rarity; it actually involves the forecasting the right-hand side variables (i.e. regressors) problem
- for the present purposes we will work with conditional forecasts





Point Forecast

• estimate the point forecast of consumption expenditure for 2005, given that $PDI_{2005} = \$31,318$ billion

$$P\hat{C}E_{2005} = b_1 + b_2 PDI_{2005}$$

= -1083.978 + 0.953768(31318)
= 28783.998
 ≈ 28784



Forecast Error

- the best mean predicted value of *PCE* in 2005 is \$28,784 billion [given $PDI_{2005} = \$31,318$]
- in table16.1 the actual value of PCE for 2005 was \$29,771 billion
- the actual value was greater than the estimated value by \$987 billion
- this is the forecast error





- since the PCE figure is an estimate, it is subject to an error
- should estimate of the forecast error we are likely to make
- if the error term u is normally distributed, then \hat{Y}_{2005} is normally distributed with mean equal to $(\beta_1 + \beta_2 X_{2005})$

$$var(\hat{Y}_{2005}) = \sigma^2 \left[\frac{1}{n} + \frac{(X_{2005} - \bar{X})^2}{\sum (X_i - \bar{X})^2} \right]$$

 \bar{X} sample mean of the X values (1960 - 2004), σ^2 variance of the error term u, n sample size





 since we do not observe the true variance u, we estimate from the sample

$$\hat{\sigma}^2 = \sum e_t^2/(n-2)$$

given the X value for 2005 we can establish a 95% confidence interval for true $E(Y_{2005})$

$$Pr[\hat{Y}_{2005} - t_{\alpha/2}se(\hat{Y}_{2005}) \le E(Y_{2005}) \le \hat{Y}_{2005} + t_{\alpha/2}se(\hat{Y}_{2005})] = 95\%$$

 $se(\hat{Y}_{2005})$ standard error obtained from $var(\hat{Y}_{2005})$ $\alpha = 5\%$



Forecast Band III

- 95% confidence interval for $E(Y_{2005})$ is (\$28,552 billion, \$29,019)
- although the single best estimate is \$28,784 billion
- we will have to compute such confidence interval for each E(Y|X) in our sample
- if we connect such confidence intervals, we obtain confidence band





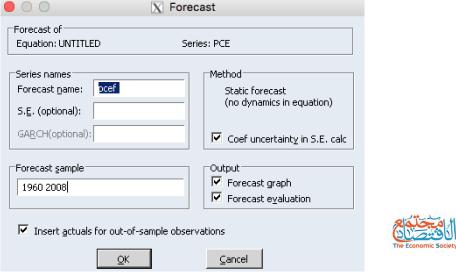
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Dependent Variable: PCE Method: Least Squares Date: 02/25/17 Time: 08:15 Sample: 1960 2004 Included observations: 45

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C PDI	-1083.978 0.953768	193.9579 0.009233	-5.588729 103.2981	0.0000 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.995986 0.995893 353.4907 5373095. -326.8829 10670.51 0.000000	Mean depend S.D. depende Akaike info cri Schwarz criter Hannan-Quin Durbin-Watso	nt var terion ion n criter.	18197.91 5515.914 14.61702 14.69731 14.64695 0.299775

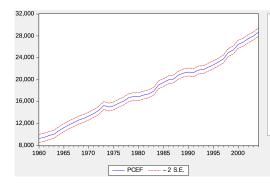






8 0 0	X Forecast
Forecast of Equation: UNTITLED	Series: P(
Series names Forecast name: ocef S.E. (optional): GARCH(optional):	Me
Forecast sample	
1960 2004	V

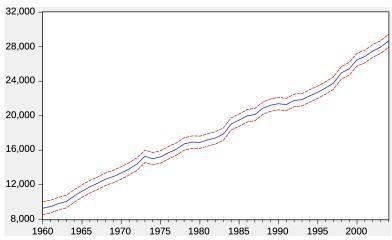




Forecast: PCEF Actual: PCE Forecast sample: 1960 2004 Included observations: 45 Root Mean Squared Error 345.5461 Mean Absolute Error 291 6878 Mean Abs. Percent Error 1.795941 Theil Inequality Coefficient 0.009095 Bias Proportion 0.000000 Variance Proportion 0.001005 Covariance Proportion 0.998995 Theil U2 Coefficient 0.715652 Symmetric MAPE 1.801425











- the variance of the estimated mean values increases as X value moves further away from its mean value
- this means that forecast error will increase as we move further away from the mean value of the regressor
- forecasting E(Y|X) for X values much greater than \bar{X} will lead to substantial forecast errors





- in the graph, measures of the quality of the forecast
 - root mean square, mean absolute error, mean absolute percentage error
 - the Theil inequality coefficient which lies between 0 and 1; the closer it is to zero, the better is the model
- these are useful if we are comparing two or more methods of forecasting





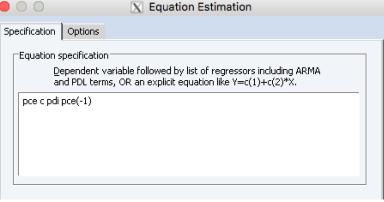
Autocorrelation

- DW statistic in regression results suggests that the error term suffers from first-order positive serial correlation
- if we can take into account serial correlation, the forecast error could be made smaller



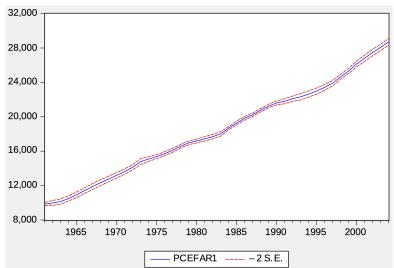


Autocorrelation





Autocorrelation







The Box-Jenkins Methodology

- BJ forecasting methodology analyses the probabilistic, or stochastic, properties of economic time series on their own
- "let the data speak for themselves"
- allows for Y_t to be explained by
 - \bullet past, lagged, values of Y_t itself, and
 - ullet current and lagged values of white noise u_t
- \bullet assumes that Y_t is stationary





Autoregressive AR Model

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \beta_2 Y_{t-2} + \dots + \beta_p Y_{t-p} + u_t$$

- u_t a white noise error term
- autoregressive model of order p, AR(p)
- ullet the value of p determined empirically using some criterion e.g., AIC





Moving Average MA Model

$$Y_t = C_0 + C_1 u_t + C_2 u_{t-1} + \cdots + C_q u_{t-q}$$

- ullet we express Y_t as a weighted, or moving, average of the current and past white noise error terms
- known as MA(q) model
- the value of q determined empirically





ARMA Model

- we can combine the AR and MA models
- called ARMA(p, q)
- p autoregressive terms, q moving average terms
- the values of p and q determined empirically





ARIMA Model

- BJ methodology assumes that Y_t is stationary, or can be made stationary by differencing
- known as ARIMA(p, d, q) model, where d denotes the number of times Y_t has to be differenced to be stationary
- in most applications d=1
- if Y_t is already stationary, then an ARIMA(p, d, q) becomes an ARMA(p,q) model





BJ Methodology in Four Steps

- identification
 - determine the appropriate values of p, d, and q
 - the main tool in this search are the correlogram and partial correlogram
- estimation
 - estimate the parameters of the chosen model
 - use OLS and sometimes use nonlinear (in parameter) estimation methods





BJ Methodology in Four Steps

- diagnostic checking
 - BJ ARIMA modelling is more an art than science
 - not absolutely sure the chosen model is the correct one
 - check the residuals are white noise otherwise start afresh (an iterative process)
- forecasting
 - the ultimate test of a successful *ARIMA* is its forecasting performance, within the sample period as well as outside the sample period





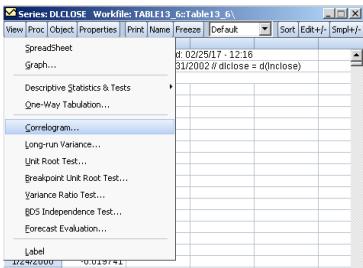
Example: IBM daily closing prices

- LCLOSE is nonstationary but DLCLOSE is stationary
- which ARMA model fits DLCLOSE
- correlgoram up to 50 lags



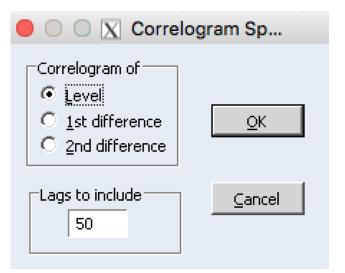


Example





Example





Example

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
dı	l di	1	-0.058	-0.058	2.3316	0.127
<u>d</u> i	l di	2	-0.069	-0.072	5.5701	0.062
III I	l ili	3	-0.010	-0.019	5.6380	0.131
1		4	0.098	0.092	12.323	0.015
ı(ı	1 1	5	-0.026	-0.017	12.802	0.025
1)1	l ilju	6	0.024	0.035	13.195	0.040
ı j ı	ļ iļi	7	0.016	0.019	13.364	0.064
ı(lı		8	-0.034	-0.039	14.192	0.077
ı ı		9	0.005	0.008	14.211	0.115
ıþı	I I	10	0.040	0.031	15.353	0.120
d i	() (i)	11	-0.059	-0.058	17.822	0.086
1(1	I I	12	-0.016	-0.011	17.993	0.116
ıþı	ļ iļi	13	0.033	0.022	18.779	0.130
ı j ı	ļ iļi	14	0.023	0.018	19.148	0.159
ı j ı	ı j	15	0.023	0.044	19.521	0.191
ıþ	ıþ	16	0.046	0.052	21.043	0.177
ıþ	iþi	17	0.050	0.058	22.789	0.156





Example: AR

		-	•	_	
Va	ariable	Coefficient	Std. Error	t-Statistic	Prob.
A A A	C NR(4) R(18) R(22) R(35)	-0.000894 0.091282 -0.080021 -0.090111 -0.058962	0.000955 0.035438 0.038859 0.036529 0.035395	-0.936224 2.575853 -2.059258 -2.466820 -1.665816	0.3495 0.0102 0.0399 0.0139 0.0962
A	R(43)	0.052808	0.045020	1.172988	0.2412





Example: AR

		<u> </u>		
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(4) AR(18) AR(22) SIGMASQ	-0.000892 0.094017 -0.086559 -0.091846 0.000676	0.000954 0.035462 0.038586 0.036372 2.52E-05	-0.935387 2.651254 -2.243281 -2.525200 26.84957	0.3499 0.0082 0.0252 0.0118 0.0000





Example MA

Coemician covariance compared doing earer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	-0.000889	0.000916	-0.970893	0.3319
MA(4)	0.085634	0.035251	2.429256	0.0154
MA(18)	-0.094356	0.040780	-2.313758	0.0210
MA(22)	-0.105041	0.037340	-2.813108	0.0050





Example ARMA (4,22)(4,22)

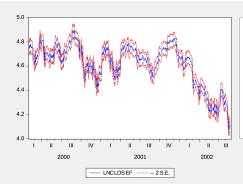
decimaters devarrance compared dering dater predictive gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	-0.000930	0.001053	-0.882798	0.3777
AR(4)	-0.294281	0.084685	-3.475008	0.0005
AR(22)	-0.604657	0.101289	-5.969596	0.0000
MA(4)	0.424877	0.085889	4.946830	0.0000
MA(22)	0.565924	0.091823	6.163217	0.0000





Forecast: ARMA (4,22)(4,22) dynamic

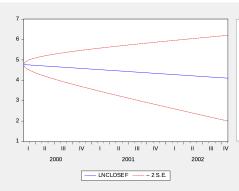


Forecast: LNCLOSEF Actual: LNCLOSE Forecast sample: 1/03/2000 10/31/2002 Adjusted sample: 2/03/2000 8/26/2002 Included observations: 668 Root Mean Squared Error 0.025958 Mean Absolute Error 0.019180 Mean Abs. Percent Error 0.418238 Theil Inequality Coefficient 0.002810 Bias Proportion 0.000004 Variance Proportion 0.001281 Covariance Proportion 0.998723 Theil U2 Coefficient 0.982139 Symmetric MAPE 0.418172





Forecast: ARMA (4,22)(4,22) static



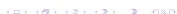
Forecast LNCLOSEF Actual: LNCLOSE Fore cast sample: 1/03/2000 10/31/2002 Adjusted sample: 2/03/2000 10/31/2002 Included observations: 716 Root Mean Squared Error 0.214929 Mean Absolute Error 0.170716 Mean Abs. Percent Error 3.661897 Theil Inequality Coefficient 0.023658 Bias Proportion 0.513064 Variance Proportion 0.002187 Covariance Proportion 0.484749 Theil U2 Coefficient 7.923331 Symmetric MAPE 3 767376

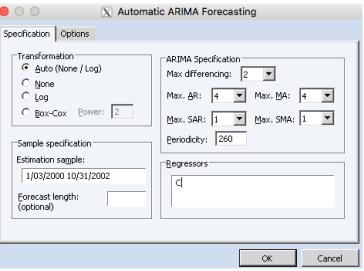




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		Generate by Equation								
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		<u>R</u> esample								
		<u>I</u> nterpolate								
		Seasonal Adjustment								
		Automatic ARIMA Forecasting								
		Forecast averaging								
		Exponential Smoothing								
		Hodrick-Prescott <u>F</u> ilter								
		Freguency Filter								
		Make <u>W</u>	hitened							
	Make <u>D</u> istribution Plot Data									











□ Summary

Automatic ARIMA Forecasting

Selected dependent variable: DLOG(LNCLOSE)

Date: 02/25/17 Time: 13:55 Sample: 1/03/2000 10/31/2002 Included observations: 686

Forecast length: 0

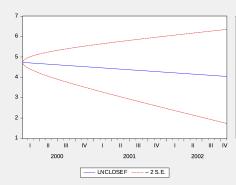
Number of estimated ARMA models: 100 Number of non-converged estimations: 0

Selected ARMA model: (0,4)(0,0)

AIC value: -6.93918150424

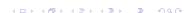


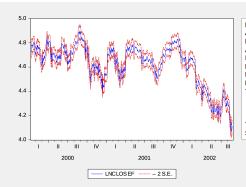




Forecast LNCLOSEF Actual: LNCLOSE Fore cast sample: 1/03/2000 10/31/2002 Adjusted sample: 1/10/2000 10/31/2002 Included observations: 734 Root Mean Squared Error 0.257808 Mean Absolute Error 0.212656 Mean Abs. Percent Error 4.569595 Theil Inequality Coefficient 0.028526 Bias Proportion 0.657006 Variance Proportion 0.003990 Covariance Proportion 0.339004 Theil U2 Coefficient 9.589423 Symmetric MAPE 4 726765







Forecast: LNCLOSEF Actual: LNCLOSE Forecast sample: 1/03/2000 10/31/2002 Adjusted sample: 1/10/2000 8/26/2002 Included observations: 686 Root Mean Squared Error 0.026157 0.019260 Mean Absolute Error Mean Abs. Percent Error 0.419634 Theil Inequality Coefficient 0.002829 Bias Proportion 0.000001 Variance Proportion 0.000946 Covariance Proportion 0.999033 Theil U2 Coefficient 0.993962 Symmetric MAPE 0.419594







